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## ON THE CONTINUITY OF INVARIANT STATISTICS

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## INTRODUCTION

The continuity of some estimates of the location and the location vector was proved in [3], Theorem 1, Hodges-Lehmann (1963), and in [4], Theorem 6.2.2, Puri-Sen (1971). All these estimates are translation invariant but the theorems do not characterize the interesting property of the estimates.

The aim of this paper is to establish theorems on the continuity of translation as well as scale invariant statistics in general, from which the above mentioned results in [3] and [4] follow.

Let us discuss the proof of the first assertion of Theorem 1 in [3] by Hodges-Lehmann (the first assertion of Theorem 6.2.2 in [4] by Puri-Sen is dealt with similarly). The authors have concluded  $P(\Delta^{**} = c) = 0$  on the basis of the Fubini theorem and of the fact that each line  $L$  of the family of all lines parallel to the direction of  $\mathbf{e} = (0, \dots, 0, 1, \dots, 1)$  with  $m$  zeros and  $n$  units intersects the set  $S = \{\Delta^{**}(X_1, \dots, X_m, Y_1, \dots, Y_n) = c\}$  in a single point which has probability zero by the assumed continuity of the cdf  $H$  of  $(X_1, \dots, X_m, Y_1, \dots, Y_n)$ . Assume  $X_1, \dots, X_m, Y_1, \dots, Y_n$  are mutually independent. The assumption makes it possible to apply the Fubini theorem, as the probability measure  $P$  is the product of its projections onto the axes  $X_1, \dots, X_m, Y_1, \dots, Y_n$ . Then the Fubini theorem could lead to the result  $P(\Delta^{**} = c) = 0$  if  $L$  were parallel to some of the axes. However,  $L$  is not so. One may try to change axes by rotating them in order to make  $L$  parallel to some new axis. But after the rotation the probability measure  $P$  is not the product of its projections onto the new axes (cf. Example 1 below) and the result  $P(\Delta^{**} = c) = 0$  cannot follow from the Fubini theorem (cf. Example 2 below).

These proofs would remain true if the following postulate were true: "the measurability of  $S$  and the fact that each  $L$  parallel to  $\mathbf{e}$  intersects  $S$  in a single point imply that  $S$  consists of a finite or denumerable family of measurable sets  $\{S_i\}$  such that for each  $S_i$  there is an axis such that each line parallel to the axis intersects  $S_i$  in at most a single point".

On the basis of this postulate Theorem 1 in Section IV would remain true if the words “absolutely continuous” were replaced by “continuous” and  $T(X, Y)$  were of the form  $T = (T_1(X_1, Y_1), \dots, T_k(X_k, Y_k))$  with the assumption that  $X^1, \dots, X^N$  are mutually independent.

Similarly, on the basis of the following postulate: “the measurability of  $S$  and the fact that each line  $L$  lying in the hyperplanes  $X_1 = x_1^0, \dots, X_m = x_m^0$ , where  $x_1^0, \dots, x_m^0$  are arbitrary real numbers, and containing the point  $(x_1^0, \dots, x_m^0, 0, \dots, 0)$  intersects  $S$  in a single point imply the assertion on  $S$  as the above postulate”, Theorem 2 in Section IV would remain true if the words “absolutely continuous” were replaced by “continuous”,  $T$  were of the form  $T = (T_1(X_1, Y_1), \dots, T_k(X_k, Y_k))$  with the assumption that  $X^1, \dots, X^N$  are mutually independent and the condition  $T_i(X, Y) = 0$  iff  $Y = 0^{(nk)}$ ,  $1 \leq i \leq k$ , were replaced by  $T_i(X_i, Y_i) = 0$  iff  $Y_i = 0^{(n)}$ ,  $1 \leq i \leq k$ .

Although the proof of the first assertion of Theorem 1 in [3] is dubious its second assertion on the absolute continuity proved in the same way remains still true, since the Lebesgue measure  $\mathcal{L}$  in  $R^N$  according to axes  $X_1, \dots, X_N$  is always the product of its projections onto the new axes  $\xi_1, \dots, \xi_N$  obtained by a rotation or by any transformation of axes  $X_1, \dots, X_N$  with the Jacobian  $J = J_1(\xi_1) \dots J_N(\xi_N)$ , and so the Fubini theorem leads to  $\mathcal{L}(A^{**} \in A) = 0$  provided  $\mathcal{L}(A) = 0$ .

In view of Lemma 2(iii) in Section II, the second assertion of Theorem 6.2.2 in [4] must be modified as follows: “the condition (6.2.8) and the absolute continuity of  $F_\alpha(x)$  for each  $\alpha = 1, \dots, n$  imply the absolute continuity of the cdf’s of each component of  $\hat{\theta}_n$ ”.

**Example 1.** Let  $P$  be a Gaussian measure with the density  $f(x, y) = (2\pi ab)^{-1} \cdot \exp\{-\frac{1}{2}(x^2/a^2 + y^2/b^2)\}$ ,  $a \neq b > 0$ . Thus  $P = P_x \times P_y$ , where  $P_x, P_y$  are projections of  $P$  onto the axes  $x$  and  $y$ . Let  $(\xi, \eta)$  be the new axes obtained by rotating  $(x, y)$  by an angle  $-\pi/4$ . Then  $P \neq P_\xi \times P_\eta$ .

**Example 2.** Suppose  $X$  and  $Y$  are dependent,  $X = Y$  a.s., and  $X$  is uniformly distributed on  $(0, 1)$ . The probability measure  $P$  generated by  $(X, Y)$  has a continuous cdf

$$\begin{aligned}
 F(x, y) &= 0 && \text{if } x \leq 0 \text{ or } y \leq 0, \\
 &x && \text{if } 0 < x < y \text{ and } 0 < x < 1, \\
 &y && \text{if } 0 < y < x \text{ and } 0 < y < 1, \\
 &1 && \text{if } x \geq 1 \text{ and } y \geq 1.
 \end{aligned}$$

Let  $S = \{x - y = 0\}$ . Each line parallel to  $y$  intersects  $S$  in a single point of probability zero by the continuity of  $F$  but  $P(S) = P(X = Y) = 1$ .

Lemmas in Section II emphasize and summarize the continuity relations between the joint cdf of a random vector and its marginal cdf’s.

## II. LEMMAS

Let  $(\Omega, \mathcal{A}, P)$  be a probability space and let  $\xi_1, \dots, \xi_k$  be random variables (a.s. finite) defined on it. Let  $F_1(x), \dots, F_k(x)$  be cdf's of  $\xi_1, \dots, \xi_k$  and let  $F(x_1, \dots, x_k)$  be the cdf of  $(\xi_1, \dots, \xi_k)$ .

**Lemma 1.** *The following assertions are equivalent:*

- (i)  $F_1(x), \dots, F_k(x)$  are continuous,
- (ii)  $F_1(x), \dots, F_k(x)$  are uniformly continuous,
- (iii)  $F(x_1, \dots, x_k)$  is continuous,
- (iv)  $F(x_1, \dots, x_k)$  is uniformly continuous.

**Proof.**

(i)  $\Leftrightarrow$  (ii) and (iv)  $\Rightarrow$  (iii) are evident.

(ii)  $\Rightarrow$  (iv): For  $x' = (x'_1, \dots, x'_k)$  and  $x'' = (x''_1, \dots, x''_k)$  let  $y' = (y'_1, \dots, y'_k)$ , where  $y'_i = \min(x'_i, x''_i)$ ,  $1 \leq i \leq k$ , and  $y'' = (y''_1, \dots, y''_k)$ , where  $y''_i = \max(x'_i, x''_i)$ ,  $1 \leq i \leq k$ . The conclusion follows from

$$\begin{aligned}
 |F(x'') - F(x')| &\leq F(y'') - F(y') = \\
 &= P\left\{ \bigcap_{i=1}^k [\xi_i \leq y'_i] \setminus \bigcap_{j=1}^k [\xi_j \leq y'_j] \right\} = \\
 &= P\left( \bigcap_i [\xi_i \leq y'_i] \cap \left( \bigcup_j [\xi_j > y'_j] \right) \right) \leq \\
 &\leq P\left\{ \bigcup_{j=1}^k \left( [\xi_j > y'_j] \cap [\xi_j \leq y'_j] \right) \right\} \leq \\
 &\leq \sum_{j=1}^k P\{y'_j < \xi_j \leq y''_j\} = \sum_{j=1}^k |F_j(x''_j) - F_j(x'_j)|.
 \end{aligned}$$

(iii)  $\Rightarrow$  (i): Suppose there is a discontinuous cdf, say  $F_1(x)$ . Let  $x_0$  be a discontinuity point of  $F_1(x)$ . Denote  $A = \xi_1^{-1}(x_0) = \{\omega : \xi_1(\omega) = x_0\} \in \mathcal{A}$ . Then

$$P(\xi_1 = x_0) = F_1(x_0) - F_1(x_0 - 0) = P(A) = p > 0.$$

Since  $F$  is continuous,

$$\begin{aligned}
 0 &= F(x_0, x_2, \dots, x_k) - F(x_0 - 0, x_2, \dots, x_k) = \\
 &= P(\xi_1 = x_0, \xi_2 \leq x_2, \dots, \xi_k \leq x_k)
 \end{aligned}$$

for any real  $x_2, \dots, x_k$ . Therefore either

$$P\{\omega \in A : \xi_2(\omega) \leq x_2, \dots, \xi_k(\omega) \leq x_k\} = 0$$

for any real  $x_2, \dots, x_k$ ,

$$\text{or } \sum_{i=2}^k P(\xi_i = \infty) \geq P(A) = p > 0,$$

which contradicts the a.s. finiteness of the random variables  $\xi_2, \dots, \xi_k$ . Q.E.D.

**Corollary 1.** *Lemma 1 remains true when  $\xi_1, \dots, \xi_k$  are random vectors of arbitrary dimensions  $n_1, \dots, n_k$ , respectively.*

**Remark 1.** The equivalence introduced in Lemma 1 is a special property of cdf's. For bounded and nondecreasing multivariate functions the equivalence does not hold in general. Let us consider

$$\begin{aligned} G(x, y) &= 0 \quad \text{if } x \leq 0 \quad \text{or } y \leq 0, \\ &(1 - e^{-y}) e^{yx} \quad \text{if } 0 < x \leq e^{-y} \quad \text{and } y > 0, \\ &1 - e^{-y} \quad \text{if } x > e^{-y} \quad \text{and } y > 0. \end{aligned}$$

Clearly,  $G(-\infty, -\infty) = 0$ ,  $G(+\infty, +\infty) = 1$  and  $G(x, y)$  is nondecreasing and continuous in  $x$  and  $y$ , but  $G(x, y)$  is not a cdf, as

$$\Delta G = G(x_2, y_2) - G(x_1, y_2) - G(x_2, y_1) + G(x_1, y_1) < 0$$

for

$$0 < x_1 < e^{-y_1} < x_2 \quad \text{and} \quad 0 < y_1 < -\ln(x_1) < y_2.$$

One has

$$\begin{aligned} G_1(x) = G(x, \infty) &= 0 \quad \text{if } x \leq 0, \\ &1 \quad \text{if } x > 0, \end{aligned}$$

i.e.,  $G_1(x)$  is not continuous and moreover,  $G(x, y)$  is not uniformly continuous, as  $G(x, 1 - \ln(x)) - G(0, 1 - \ln(x)) > 1 - e^{-1} > \frac{1}{2}$  for any  $x$ ,  $0 < x < 1$ .

**Lemma 2.**

- (i) *If  $F(x_1, \dots, x_k)$  is absolutely continuous, then  $F_1(x), \dots, F_k(x)$  are so as well.*
- (ii) *If  $F_1(x), \dots, F_k(x)$  are absolutely continuous and  $\xi_1, \dots, \xi_k$  are mutually independent, then  $F(x_1, \dots, x_k)$  is so as well.*
- (iii) *If  $F_1(x), \dots, F_k(x)$  are absolutely continuous, then  $F(x_1, \dots, x_k)$  is not generally so but it is continuous even if  $\xi_1, \dots, \xi_k$  are mutually dependent.*

**Proof.**

(i) and (ii) are well-known.

(iii) The continuity of  $F$  follows from Lemma 1. In order to prove that  $F$  is generally not absolutely continuous it is sufficient to form a counterexample. Let  $k = 2$ . Let  $\xi_1 = \xi_2$  a.s. and let each of them be uniformly distributed on  $(0, 1)$ . From

Example 2 one has clearly  $\partial^2 F / (\partial x_1 \partial x_2) = 0$  a.e. with respect to the Lebesgue measure  $\mathcal{L}$  in  $R^2$ , and  $\iint_s dF(x_1, x_2) = 1$  where  $s = \{(x, x), 0 < x < 1\}$  with  $\mathcal{L}(s) = 0$ . It means that  $F(x_1, x_2)$  is not absolutely continuous (but is continuous), while  $F_1(x)$  and  $F_2(x)$  are so. Q.E.D.

Another example which is not so special is the following

Example 3. Let  $\xi_1$  be uniformly distributed on  $(0, 1)$  and  $P(\xi_2 \in (0, 1)) = 1$ . Let  $\xi_2 = \xi_1$  for  $0 < \xi_1 \leq \frac{1}{2}$ , and let  $\xi_2$  be uniformly distributed on  $(\frac{1}{2}, 1)$  for  $\frac{1}{2} < \xi_1 \leq 1$ , i.e. the conditional cdf of  $\xi_2$  for a given  $\xi_1$  is of the form

$$F_2(x_2 | x_1) = 0 \text{ if } 0 < x_1 \leq \frac{1}{2} \text{ and } x_2 < x_1, \\ 2x_2 - 1 \text{ if } \frac{1}{2} < x_1 < 1 \text{ and } \frac{1}{2} < x_2 < 1, \\ 1 \text{ if } 0 < x_1 \leq \frac{1}{2} \text{ and } x_2 \geq x_1, \text{ or } \frac{1}{2} < x_1 < 1 \text{ and } x_2 \geq 1.$$

Therefore

$$F(x_1, x_2) = 0 \text{ if } x_1 \leq 0 \text{ or } x_2 \leq 0, \\ x_1 \text{ if } 0 < x_1 \leq \frac{1}{2} \text{ and } x_2 \geq x_1, \text{ or } \frac{1}{2} < x_1 < 1 \text{ and } x_2 \geq 1, \\ x_2 \text{ if } 0 < x_2 \leq \frac{1}{2} \text{ and } x_1 \geq x_2, \text{ or } \frac{1}{2} < x_2 < 1 \text{ and } x_1 \geq 1, \\ \frac{1}{2} + 2(x_1 - \frac{1}{2})(x_2 - \frac{1}{2}) \text{ if } \frac{1}{2} < x_1 < 1 \text{ and } \frac{1}{2} < x_2 < 1, \\ 1 \text{ if } x_1 \geq 1 \text{ and } x_2 \geq 1.$$

Thus we obtain

$$F_2(x_2) = 0 \text{ if } x_2 \leq 0, \\ x_2 \text{ if } 0 < x_2 < 1, \\ 1 \text{ if } x_2 \geq 1,$$

i.e.  $\xi_2$  is also uniformly distributed on  $(0, 1)$  as  $\xi_1$ , while

$$\frac{\partial^2 F}{\partial x_1 \partial x_2} = \begin{cases} 2 & \text{if } \frac{1}{2} < x_1 < 1 \text{ and } \frac{1}{2} < x_1 < 1, \\ 0 & \text{otherwise,} \end{cases} \\ \iint_{R^2} \frac{\partial^2 F}{\partial x_1 \partial x_2} dx_1 dx_2 = \frac{1}{2} \text{ and } \iint_s dF(x_1, x_2) = \frac{1}{2},$$

where  $s = \{(x, x), 0 < x < \frac{1}{2}\}$  with  $\mathcal{L}(s) = 0$ , i.e.  $F$  is not absolutely continuous (but is continuous).

**Corollary 2.** Lemma 2 can be generalized for the case of  $\xi_1, \dots, \xi_k$  being random vectors.

### III. NOTATION AND DEFINITIONS

Let  $X^j = (X_1^j, \dots, X_k^j)$ ,  $1 \leq j \leq N$  be  $k$ -dimensional random vectors. Let  $X = (X^1, \dots, X^m)$ ,  $Y = (X^{m+1}, \dots, X^{m+n})$ ,  $Z = (X, Y)$ , where  $m \geq 0$ ,  $n \geq 1$ ,  $m + n = N$ . For the case  $m = 0$ ,  $Y = Z$ , let  $x, y, z, \dots$  be representations of  $X, Y, Z, \dots$ , respectively. Denote  $Z_i = (X_i, Y_i)$ ,  $1 \leq i \leq k$ , where

$$X_i = (X_i^1, \dots, X_i^m), \quad Y_i = (X_i^{m+1}, \dots, X_i^{m+n}).$$

Let  $T = T(Z) = T(X, Y) = (T_1(X, Y), \dots, T_k(X, Y))$  be a  $k$ -dimensional statistic. For  $a = (a_1, \dots, a_s) \in R^s$ ,  $b = (b_1, \dots, b_s) \in R^s$ , let  $a^{(p)}$  stand for  $(a, \dots, a) \in R^{ps}$  and  $a * b = (a_1 b_1, \dots, a_s b_s)$ .

**Definition 1.** The statistic  $T(X, Y)$  is said to be translation invariant iff

$$(1) \quad T(X, Y + b^{(n)}) = T(X, Y) + b \quad \text{for all } b = (b_1, \dots, b_k) \in R^k.$$

**Definition 2.** The statistic  $T(X, Y)$  is said to be scale invariant of the first type or of the second type iff

$$(2) \quad T(X, tY) = t T(X, Y) \quad \text{for all } t \in R^1, \text{ or}$$

$$(3) \quad T(X, a^{(n)} * Y) = a * T(X, Y) \quad \text{for all } a = (a_1, \dots, a_k) \in R^k,$$

respectively.

**Definition 3.** The statistic  $T$  is said to be linear invariant of the first type or of the second type iff it is translation invariant as well as scale invariant of the first type or of the second type, i.e. iff

$$(4) \quad T(X, tY + b^{(n)}) = t T(X, Y) + b \quad \text{for all } t \in R^1 \text{ and all } b \in R^k, \text{ or}$$

$$(5) \quad T(X, [a^{(n)} * Y] + b^{(n)}) = [a * T(X, Y)] + b \quad \text{for all } a, b \in R^k,$$

respectively.

**Remark 2.** In some cases (if necessary) Definitions 2 and 3 may be modified in the following way: (2), (3), (4), and (5) are replaced by

$$(2') \quad T(tX, tY) = t T(X, Y) \quad \text{for all } t \in R^1,$$

$$(3') \quad T(a^{(m)} * X, a^{(n)} * Y) = a * T(X, Y) \quad \text{for all } a \in R^k,$$

$$(4') \quad T(tX, tY + b^{(n)}) = t T(X, Y) + b \quad \text{for all } t \in R^1 \text{ and all } a \in R^k,$$

$$(5') \quad T(a^{(m)} * X, [a^{(n)} * Y] + b^{(n)}) = [a * T(X, Y)] + b \quad \text{for all } a, b \in R^k.$$

Note that Definitions 2 and 3 of the second type are stronger than those of the first type. Statistics satisfying one or all the Definitions would be formulated in estimating location parameters, see e.g. Hodges-Lehmann (1963) [3], Puri-Sen (1971) [4], Bickel-Lehmann (1975) [1], [2], ...

**Remark 3.** If  $T(X, Y) = (T_1(X_1, Y_1), \dots, T_k(X_k, Y_k))$ ,  $T$  is scale or linear invariant of the second type iff it is scale or linear invariant of the first type, respectively, as obtained easily from the Definitions.

#### IV. THEOREMS

Let us keep the notation of Section III.

**Theorem 1.** *Let  $T(X, Y) = (T_1(X, Y), \dots, T_k(X, Y))$  be translation invariant. Then the cdf's of  $T_1, \dots, T_k$  are absolutely continuous provided the cdf  $F(x, y)$  of  $(X, Y)$  is so.*

*Proof.* For  $i, 1 \leq i \leq k$  and  $A \subset R^1$  with  $\mathcal{L}(A) = 0$ , put  $R_i^A = \{(x, y) \in R^{Nk} : T_i(x, y) \in A\}$ .  $R_i^A$  is measurable. Consider in  $R^{Nk}$  the family of all lines parallel to the direction of the vector  $I = (0^{(mk)}, 1^{(nk)})$ :

$$\mathcal{L}_I = \{L(x, y) = \{(x, y + t^{(nk)}), t \in R^1\}, (x, y) \in R^{Nk}\}.$$

For each  $L = L(x^0, y^0) \in \mathcal{L}_I$ , one has

$$[(x, y) \in L(x^0, y^0) \cap R_i^A] \Leftrightarrow [(x, y) = (x^0, y^0 + [c - T_i(x^0, y^0)]^{(nk)}), c \in A]$$

This means that each  $L \in \mathcal{L}_I$  intersects  $R_i^A$  in a set equivalent to  $A$ , therefore the set has the Lebesgue measure zero. Then  $\mathcal{L}(R_i^A) = 0$ , by the Fubini theorem applied to the Lebesgue measure  $\mathcal{L}$  in  $R^{Nk}$ . It follows from the absolute continuity of  $F(x, y)$  that  $P(T_i \in A) = \int_{R_i^A} \dots \int dF(x, y) = 0$  for all  $A \subset R^1$  with  $\mathcal{L}(A) = 0$ , i.e. the cdf of  $T_i$  is absolutely continuous. Q.E.D.

**Remark 4.** In view of Lemma 2 (iii) one cannot obtain the absolute continuity of the joint cdf of  $T = (T_1, \dots, T_k)$  under the assumptions of Theorem 1. The same argument explains why the second assertion of Theorem 6.2.2 in [4] mentioned in Section I is dubious.

**Corollary 3.** *The result of Theorem 1 holds for a statistic  $T$  which is linear invariant of the first type or of the second type.*

**Theorem 2.** *Let  $T = (T_1(X, Y), \dots, T_k(X, Y))$  be scale invariant of the first type or of the second type and such that  $T_i(x, y) = 0$  iff  $y = 0^{(nk)}$ ,  $1 \leq i \leq k$ . Let  $F(x, y)$  be absolutely continuous. Then the cdf's of  $T_1, \dots, T_k$  are so as well.*



Proof. Theorem 2 is proved similarly as Theorem 1, with  $\mathcal{L}_I$  replaced by

$$\mathcal{L}'_I = \{L(x, y) = \{(x, ty) : t \in R^1\}, (x, y) \in R^{Nk}, y \neq 0^{(nk)}\}.$$

The intersection of  $L(x^0, y^0)$  with  $R^A_i, A \subset R^1, \mathcal{L}(A) = 0$  is  $\{(x^0, [c/(T(x^0, y^0))] y^0), c \in A$ , which has the Lebesgue measure zero. The rest of the proof follows as in Theorem 1. Q.E.D.

Remark 5. Theorem 2 remains true for  $T$  which is scale invariant in the sense of the definition modified as in Remark 2 and such that  $T_i(x, y) = 0$  iff  $(x, y) = 0^{(Nk)}$ . In order to prove it let us put  $\mathcal{L}'_I = \{L(x, y) = \{t(x, y), t \in R^1\}, (x, y) \in R^{Nk} \setminus \{0^{(Nk)}\}\}$ . Then the intersection of  $L(x^0, y^0) \in \mathcal{L}'_I$  with  $R^A_i, A \subset R^1, \mathcal{L}(A) = 0$  is  $\{[c/(T_i(x^0, y^0))] (x^0, y^0), c \in A\}$  of the Lebesgue measure zero.

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#### Souhrn

### SPOJITOST INVARIANTNÍCH STATISTIK

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Cílem článku je dokázat obecné věty o absolutní spojitosti statistik invariantních vzhledem ke změně polohy a měřítka, z nichž plynou příbuzné výsledky Hodgese-Lehmannova a Puri-Sena. Vyšetřuje se také vztah mezi spojitostí sdružené distribuční funkce náhodného vektoru a jeho marginálních distribučních funkcí.

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