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## How non-symmetric can a copula be?

ERICH PETER KLEMENT, RADKO MESIAR

*Abstract.* A two-place function measuring the degree of non-symmetry for (quasi-)copulas is considered. We construct copulas which are maximally non-symmetric on certain subsets of the unit square. It is shown that there is no copula (and no quasi-copula) which is maximally non-symmetric on the whole unit square.

*Keywords:* copula, quasi-copula, symmetry, opposite diagonal

*Classification:* Primary 62H05; Secondary 62E10

### 1. Introduction

Copulas (first mentioned in [11], for an excellent survey see [9]) and quasi-copulas (introduced in [1] and conveniently characterized in [4]) play a key role in the analysis of bivariate distribution functions with given marginals. The basic result in this context is Sklar's Theorem ([11], [12]) showing that the joint distribution of a random vector and the corresponding marginal distributions are linked by some copula: if  $(X, Y)$  is a random vector,  $F_X, F_Y: [-\infty, \infty] \rightarrow [0, 1]$  are its marginal distribution functions, then  $H_{XY}: [-\infty, \infty]^2 \rightarrow [0, 1]$  is a joint distribution of  $(X, Y)$  if and only if there is a two-dimensional copula  $C_{XY}$  such that for all  $(x, y) \in [-\infty, \infty]^2$  we have

$$H_{XY}(x, y) = C_{XY}(F_X(x), F_Y(y)).$$

Moreover, if  $F_X$  and  $F_Y$  are continuous then  $C_{XY}$  is unique, otherwise  $C_{XY}$  is uniquely determined only on  $\text{Ran}(F_X) \times \text{Ran}(F_Y)$ .

Recall that a (*two-dimensional*) *copula* is a function  $C: [0, 1]^2 \rightarrow [0, 1]$  such that  $C(0, x) = C(x, 0) = 0$  and  $C(1, x) = C(x, 1) = x$  for all  $x \in [0, 1]$ , and  $C$  is 2-increasing, i.e., for all  $x_1, x_2, y_1, y_2 \in [0, 1]$  with  $x_1 \leq x_2$  and  $y_1 \leq y_2$  we have

$$C(x_1, y_1) + C(x_2, y_2) \geq C(x_1, y_2) + C(x_2, y_1).$$

A *quasi-copula* is a function  $Q: [0, 1]^2 \rightarrow [0, 1]$  such that  $Q(0, x) = Q(x, 0) = 0$  and  $Q(1, x) = Q(x, 1) = x$  for all  $x \in [0, 1]$ ,  $Q$  is non-decreasing (in each component), and  $Q$  is 1-Lipschitz, i.e., for all  $x_1, x_2, y_1, y_2 \in [0, 1]$

$$|Q(x_1, y_1) - Q(x_2, y_2)| \leq |x_1 - x_2| + |y_1 - y_2|.$$

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Obviously, each copula is a quasi-copula, but not vice versa. Each copula  $C$  satisfies

$$(1.1) \quad W \leq C \leq M,$$

where the Fréchet-Hoeffding lower and upper bounds  $W$  and  $M$  are given by  $W(x, y) = \max(x + y - 1, 0)$  and  $M(x, y) = \min(x, y)$ , respectively, and the same holds for quasi-copulas.

In general, a copula is neither symmetric (commutative) nor associative (see [8]), and it is well-known that each associative copula is also symmetric and, consequently, a (continuous) triangular norm [6], [10] (again the converse does not necessarily hold).

There is a close relationship between symmetric copulas and interchangeable random variables  $X$  and  $Y$  (where the random vectors  $(X, Y)$  and  $(Y, X)$  are identically distributed). Clearly, two interchangeable random variables  $X$  and  $Y$  must be identically distributed, i.e., have a common univariate distribution function, and for identically distributed random variables  $X$  and  $Y$  their interchangeability is equivalent to the symmetry of their copula  $C_{XY}$  (see [9, Theorem 2.7.4]).

As a consequence, for exchangeable random variables  $X$  and  $Y$  with copula  $C$ , the symmetry of  $C$  implies  $C(y, x) = C(x, y)$ . In general (i.e., for non-exchangeable random variables  $X$  and  $Y$ ) this is no more true, but any estimate of the value  $C(y, x)$  by means of  $C(x, y)$  will be helpful when modelling bivariate statistical data, especially in order to exclude irrelevant models.

Therefore, we are interested in “how non-symmetric” a copula can be, and we construct copulas which are “maximally” non-symmetric on certain distinguished subsets of the unit square. Finally we show that no copula (and no quasi-copula) can be “maximally” non-symmetric on the whole unit square.

## 2. Degree of non-symmetry

Given a copula  $C$ , the function  $d_C: [0, 1]^2 \rightarrow [0, 1]$  defined by

$$d_C(x, y) = |C(x, y) - C(y, x)|$$

provides a “measure” of its non-symmetry at each point of the unit square  $[0, 1]^2$ , and its Chebyshev norm  $\|d_C\|_\infty$  given by

$$\|d_C\|_\infty = \sup\{d_C(x, y) \mid (x, y) \in [0, 1]^2\}$$

can be viewed as the *degree of non-symmetry* of  $C$ . Obviously, for each copula  $C$  the function  $d_C$  vanishes on the boundary as well as on the diagonal  $\{(x, x) \mid x \in [0, 1]\}$  of  $[0, 1]^2$ . Also, a copula  $C$  is symmetric if and only if  $\|d_C\|_\infty = 0$ .

**Example 2.1.** The copula  $C$  given by  $C(x, y) = xy - x^3y(1-x)(1-y)$  is non-symmetric, and we obtain  $d_C(x, y) = xy(1-x)(1-y)|x^2 - y^2|$ . A simple computation then yields  $\|d_C\|_\infty = d_C(0.3418922, 0.7768102) = 0.0189801$ .

In order to find out the maximal degree of non-symmetry of copulas consider the function  $d^*: [0, 1]^2 \rightarrow [0, 1]$  defined by

$$d^* = \sup\{d_C \mid C \text{ is a copula}\}.$$

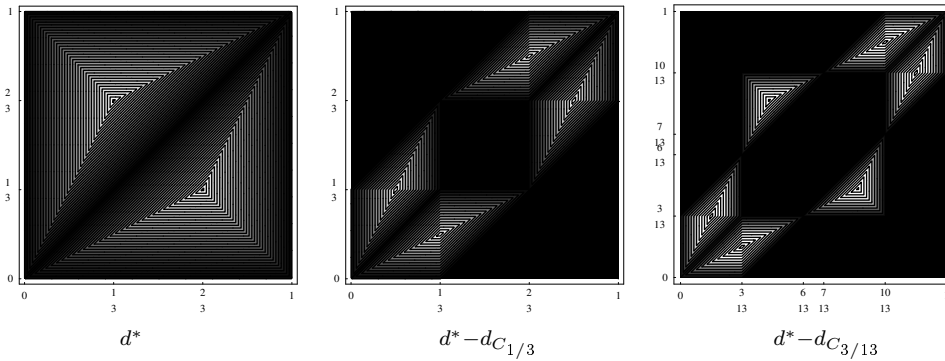


FIGURE 1: Maximal non-symmetry

We now compute the function  $d^*$  (see Figure 1 left) and show that for each point  $(x_0, y_0) \in [0, 1]^2$  we can find a copula  $C$  such that  $d_C$  and  $d^*$  coincide on two straight line segments containing the points  $(x_0, y_0)$  and  $(y_0, x_0)$ .

**Proposition 2.2.**

- (i) For all  $(x, y) \in [0, 1]^2$  we have  $d^*(x, y) = \min(|x - y|, x, y, 1 - x, 1 - y)$ .
- (ii) For each  $\lambda \in [0, 1]$  the function  $C_\lambda: [0, 1]^2 \rightarrow [0, 1]$  given by

$$C_\lambda(x, y) = \max(M(x, y - \lambda), W(x, y))$$

is a copula such that we have  $d_{C_\lambda}(x, y) = d^*(x, y)$  for all  $(x, y) \in [0, 1]^2$  with  $|x - y| = \lambda$ .

PROOF: Let  $C$  be a copula and assume, without loss of generality,  $x \leq y$  and  $C(x, y) \leq C(y, x)$ . Then the monotonicity of  $C$  yields  $C(x, y) \leq C(y, x) \leq C(y, y)$  which, together with (1.1) and the fact that  $C$  is 1-Lipschitz, implies  $d_C(x, y) \leq \min(|x - y|, M(x, y) - W(x, y))$ . A simple computation shows that the latter expression coincides with  $\min(|x - y|, x, y, 1 - x, 1 - y)$ , i.e., for all  $(x, y) \in [0, 1]^2$

$$d_C(x, y) \leq \min(|x - y|, x, y, 1 - x, 1 - y).$$

Now fix an arbitrary point  $(x_0, y_0) \in [0, 1]^2$  and put  $\lambda = |x_0 - y_0|$ . If we can show that  $C_\lambda$  in (ii) is a copula satisfying

$$(2.1) \quad d_{C_\lambda}(x_0, y_0) = \min(\lambda, x_0, y_0, 1 - x_0, 1 - y_0)$$

this will complete the proof of (i).

Since  $C_\lambda$  is a shuffle of  $M$  it is a copula (see [9]). Note that for each  $(x, y) \in [0, 1]^2$

$$d_{C_\lambda}(x, y) = \min(\max(\min(x - \lambda, y, 1 - x, 1 - \lambda - y), \min(y - \lambda, x, 1 - y, 1 - \lambda - x), 0), |x - y|, \lambda).$$

Then the verification of (2.1) is a matter of simple but tedious checking of all possible cases. Since  $\lambda$  only depends on  $|x_0 - y_0|$ , the proof of (ii) is complete, too.  $\square$

An immediate consequence of Proposition 2.2 is the following:

**Corollary 2.3.** *For each copula  $C$  and each  $(x, y) \in [0, 1]^2$  we have:*

$$C(y, x) \in [\max(W(y, x), C(x, y) - |x - y|), \min(M(y, x), C(x, y) + |x - y|)].$$

Observe that the estimate for  $C(y, x)$  in Corollary 2.3 is better than the estimate derived from the Fréchet-Hoeffding bounds  $W$  and  $M$ : if for a copula  $C$  we have  $C(0.5, 0.6) = 0.3$  then the Fréchet-Hoeffding bounds imply  $C(0.6, 0.5) \in [0.1, 0.5]$ , whereas Corollary 2.3 tells us  $C(0.6, 0.5) \in [0.2, 0.4]$ .

Although copulas form a proper subclass of the class of quasi-copulas, the fact that we did not need the 2-increasingness of copulas implies:

**Corollary 2.4.** *We also have  $d^* = \sup\{d_Q \mid Q \text{ is a quasi-copula}\}$ .*

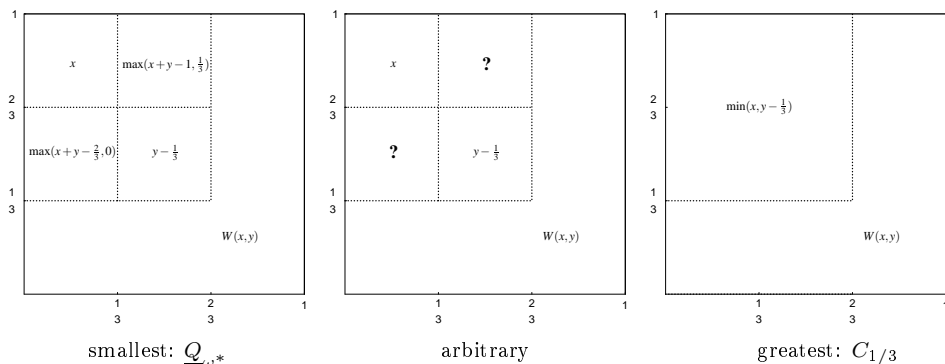


FIGURE 2: Copulas with opposite diagonal  $\omega^*$

Some straightforward calculations show that the maximal value of  $d^*$  equals  $\frac{1}{3}$  and that there is indeed a copula, namely,  $C_{1/3}$  (see Figure 2 right) such that  $d_{C_{1/3}}$  attains this maximal value in the points  $(\frac{1}{3}, \frac{2}{3})$  and  $(\frac{2}{3}, \frac{1}{3})$  (see Figure 1 center):

**Corollary 2.5.**

(i) For each  $\lambda \in [0, \frac{1}{3}]$  we have

$$\|d_{C_\lambda}\|_\infty = d_{C_\lambda}(\lambda, 1 - \lambda) = \lambda.$$

(ii) In particular, we have

$$\|d^*\|_\infty = d^*(\frac{1}{3}, \frac{2}{3}) = \frac{1}{3} = d_{C_{1/3}}(\frac{1}{3}, \frac{2}{3}) = \|d_{C_{1/3}}\|_\infty.$$

**Example 2.6.** From the proof of Proposition 2.2 it follows that in the class  $\mathcal{S}_M$  of shuffles of  $M$  (see [9]) for each  $(x_0, y_0) \in [0, 1]^2$  we can find a copula  $C \in \mathcal{S}_M$  such that  $d_C(x_0, y_0) = d^*(x_0, y_0)$ . For other well-known classes of copulas this does not hold:

- (i) Evidently, for each associative (and, consequently, for each Archimedean) copula  $C$  the function  $d_C$  vanishes on the whole unit square  $[0, 1]^2$ .
- (ii) In the class  $\mathcal{A}$  of maximum attractors [2] (compare also [3], [7]) we obtain

$$\sup\{\|d_C\|_\infty \mid C \in \mathcal{A}\} = \frac{1}{5} \cdot (\frac{4}{5})^4.$$

This extremal value is attained in the points  $(\frac{1}{5}, \frac{2}{3})$  and  $(\frac{2}{5}, \frac{1}{3})$  by the function  $d_{C_A}$ , where the maximum attractor  $C_A$  is given by

$$C_A(x, y) = (xy)^{A(\frac{\log x}{\log(xy)})}$$

and the dependence function  $A: [0, 1] \rightarrow [0, 1]$  by

$$A(x) = \begin{cases} 1 - x & \text{if } x \in [0, \frac{1}{3}], \\ \frac{x+1}{2} & \text{otherwise.} \end{cases}$$

**Example 2.7.** Although for the copula  $C_{1/3}$  we know that  $d_{C_{1/3}}$  attains the maximal value of  $d^*$  in  $(\frac{1}{3}, \frac{2}{3})$ , there are other members of the family  $(C_\lambda)_{\lambda \in [0, 1]}$  such that the area of the subset of  $[0, 1]^2$  on which  $d_{C_\lambda}$  and  $d^*$  coincide is greater. In general, for  $\lambda \in [0, 1]$  the area of the subset of  $[0, 1]^2$  on which  $d_{C_\lambda}$  and  $d^*$  coincide equals  $(1 - \lambda)^2 + (\max(1 - 2\lambda, 0))^2 - 2(\max(1 - 3\lambda, 0))^2$ , assuming its maximal value  $\frac{9}{13}$  for  $\lambda = \frac{3}{13}$  (see Figure 1 right).

### 3. Non-symmetry and opposite diagonal section

A closer look at the copula  $C_{1/3}$  shows that the functions  $d_{C_{1/3}}$  and  $d^*$  coincide on  $[\frac{1}{3}, \frac{2}{3}]^2 \cup \{(x, y) \in [0, 1]^2 \mid |x - y| \geq \frac{1}{3}\}$  (see Figure 1 center). This means, in particular, that we have  $d_{C_{1/3}}(x, 1 - x) = d^*(x, 1 - x)$  for all  $x \in [0, 1]$ , i.e.,  $C_{1/3}$  is “maximally non-symmetric” on the whole opposite diagonal  $\{(x, 1 - x) \mid x \in [0, 1]\}$  of the unit square  $[0, 1]^2$  (note that  $C_{1/3}$  is the only copula in the family  $(C_\lambda)_{\lambda \in [0, 1]}$  with this property).

From [5] we know that, for a given copula  $C$ , its opposite diagonal section  $\omega_C: [0, 1] \rightarrow [0, 1]$  defined by  $\omega_C(x) = C(x, 1 - x)$  must be a 1-Lipschitz function satisfying  $\omega_C(0) = \omega_C(1) = 0$ .

Therefore, if for some copula  $C$  we require  $d_C(x, 1 - x) = d^*(x, 1 - x)$  for all  $x \in [0, 1]$ , the only possibilities are either  $\omega_C = \omega^*$  or  $\omega_C = \omega_1$ , where the functions  $\omega^*, \omega_1: [0, 1] \rightarrow [0, 1]$  are given by

$$\begin{aligned} \omega^*(x) &= \max(\min(x, \frac{2}{3} - x), 0), \\ \omega_1(x) &= \max(\min(1 - x, x - \frac{1}{3}), 0). \end{aligned}$$

However, if for some (necessarily non-symmetric) copula  $C$  we have  $\omega_C = \omega^*$  then for the copula  $C_1$  defined by  $C_1(x, y) = C(y, x)$  we have  $\omega_{C_1} = \omega_1$ . This means that we can restrict our considerations to copulas  $C$  satisfying  $\omega_C = \omega^*$ .

From [5, Proposition 7.3] it follows that  $C_{1/3}$  is just the greatest copula with opposite diagonal section  $\omega^*$ . Moreover, because of [5, Proposition 6.5(ii)] the smallest quasi-copula  $\underline{Q}_{\omega^*}$  with opposite diagonal section  $\omega^*$  is given by

$$\underline{Q}_{\omega^*}(x, y) = \begin{cases} x & \text{if } (x, y) \in [0, \frac{1}{3}] \times [\frac{2}{3}, 1], \\ \max(x + y - \frac{2}{3}, 0) & \text{if } (x, y) \in [0, \frac{1}{3}] \times [\frac{1}{3}, \frac{2}{3}[ , \\ \max(x + y - 1, \frac{1}{3}) & \text{if } (x, y) \in ]\frac{1}{3}, \frac{2}{3}] \times [\frac{2}{3}, 1], \\ y - \frac{1}{3} & \text{if } (x, y) \in ]\frac{1}{3}, \frac{2}{3}] \times [\frac{1}{3}, \frac{2}{3}[ , \\ W(x, y) & \text{otherwise.} \end{cases}$$

For our special opposite diagonal section  $\omega^*$ , the quasi-copula  $\underline{Q}_{\omega^*}$  turns out to be a copula since it is again a shuffle of  $M$  (see Figure 2 left).

With these preliminary considerations, we are able to show:

**Proposition 3.1.** *There is no copula  $C$  such that  $d_C = d^*$ .*

PROOF: Suppose that  $C$  is a copula such that  $d_C = d^*$ . Then, in particular,  $d_C$  and  $d^*$  must coincide on the opposite diagonal, i.e., we must have either  $\omega_C = \omega^*$  or  $\omega_C = \omega_1$ . Assume without loss of generality that  $\omega_C = \omega^*$ . Since  $\underline{Q}_{\omega^*}$  and  $C_{1/3}$  are the smallest and greatest copula with opposite diagonal section  $\omega^*$ , it

follows immediately that each copula  $C$  with  $\omega_C = \omega^*$  coincides with  $\underline{Q}_{\omega^*}$  and  $C_{1/3}$  on

$$[0, 1]^2 \setminus \left( ([0, \frac{1}{3}[\times] \frac{1}{3}, \frac{2}{3}[) \cup ([\frac{1}{3}, \frac{2}{3}[\times] \frac{2}{3}, 1[) \right)$$

(see Figure 2 center — the question marks indicate the regions where  $C$  is not uniquely determined by the lower and upper bounds  $\underline{Q}_{\omega^*}$  and  $C_{1/3}$ ). As a consequence,  $C$  coincides with the symmetric copula  $W$  on the set  $[0, \frac{1}{3}]^2 \cup [\frac{2}{3}, 1]^2$ , implying that  $d_C$  vanishes on this set. Since  $d^*$  vanishes only on the boundary and the diagonal of  $[0, 1]^2$  this shows that for no copula  $C$  the equality  $d_C = d^*$  can hold.  $\square$

Since again the 2-increasingness of copulas was not used in our argument, we also have shown:

**Corollary 3.2.** *There is no quasi-copula  $Q$  such that  $d_Q = d^*$ .*

**Example 3.3.** Clearly, for each symmetric (quasi-)copula  $C$  the value  $\|d^* - d_C\|_\infty$  attains its maximum  $\frac{1}{3}$ . For the family  $(C_\lambda)_{\lambda \in [0,1]}$  of copulas considered in Proposition 2.2(ii) we obtain  $\|d^* - d_{C_\lambda}\|_\infty = \min(\max(\frac{1}{3} - \lambda, \frac{\lambda}{2}), \frac{1}{3})$ . This value is minimal for  $\lambda = \frac{2}{9}$ , and we get  $\|d^* - d_{C_{2/9}}\|_\infty = \frac{1}{9}$ . Observe, however, that also for  $\lambda \in [\frac{2}{3}, 1[$  we get the maximal value  $\|d^* - d_{C_\lambda}\|_\infty = \frac{1}{3}$ , although the corresponding copulas  $C_\lambda$  are non-symmetric.

**Note added in proof:** Similar results were obtained independently by R.B. Nelsen (Extremes of nonexchangeability, Statist. Papers, to appear).

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